

From last week: Majority element.

Given a sequence $S = x_1, x_2, \dots, x_n$ of numbers

Q: is there an index $i \in \{1, 2, \dots, n\}$ s.t.
 x_i occurs $> \frac{n}{2}$ times in S ?

Easy to solve in time (# comparison) $O(n \log n)$
 using sorting

Randomized majority (S, m)

Repeat m times

A: select random $i \in [n]$
 if x_i is the majority, then return 'true' $O(n)$
 end
 return 'false'

If A returns 'true' then there is a majority, so the answer is correct

If A returns 'false', then there could still be a majority element, however the probability^F that A is wrong is at most $(\frac{1}{2})^m$ ($p < (\frac{1}{2})^m$)

Why? in each round, the probability of not selecting the majority^x (assuming there is one) is less than $\frac{1}{2}$

$x_j = x$	$x_j \neq x$
$> \frac{1}{2}n$	$< \frac{1}{2}n$

Is this enough to be useful?

Yes because we can make p arbitrarily small:

m	10	20	30
p	2^{-10}	2^{-20}	2^{-30}

A is an example of a Monte-Carlo alg
these have two outputs 'true'/'false'
one of which is always correct and the
other only wrong with a fixed prob.

Ex 15 Quality control

know : If container was checked
 then no bad chips
 otherwise 10% of chips are bad

$$p=0$$

checked

$$p=\frac{1}{10}$$

not checked

 $p = \text{prob of bad chip}$

Monte-carlo strategy (m)

Repeat m times

B

pick random chip

if bad then return 'false'

end

return 'true'

Observations

B is always correct if it returns 'false'

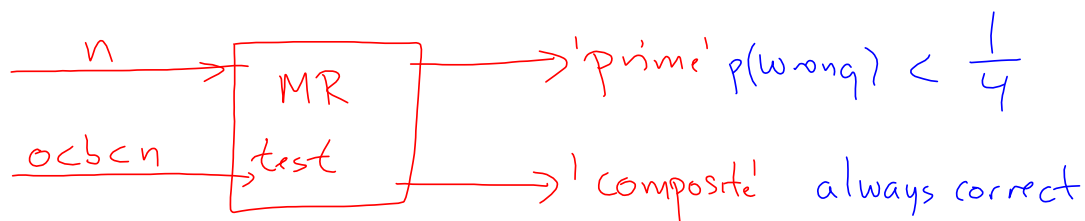
B is wrong (by answering 'true')

with probability at most $(1-p)^m$ in this case $p = \frac{1}{10}$ so wrong with

$$\text{prob} < \left(\frac{9}{10}\right)^m$$

if $m = 136$ then error with prob $< 10^{-6}$

Prime testing (Miller-Rabin) MR



MC algorithm (n, m)

repeat m times

m

select random $b \in \{1, 2, \dots, n-1\}$

do MR-test(n, b)

if 'composite' then return 'composite'

end

return 'prime'

M is wrong with probability at most $\left(\frac{1}{4}\right)^m$

Ex 13 Birthday problem.

Q: what is the minimum # of persons in a room so that probⁿ that two have the same birthday is at least $1/2$?

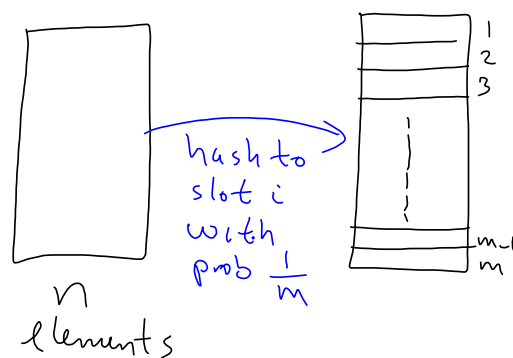
assumption . 1 year = 366 days
 . all birthdays are equally likely

$q = (1-p)$ is prob. of all different
 (we want n s.t. $q < \frac{1}{2}$)

366 dates for person 1
 365 - - - - - 2
 ⋮
 367-j - - - - - j

$$q = \frac{366}{366} \cdot \frac{365}{366} \cdot \dots \cdot \frac{(367-n)}{366} \approx \left. \begin{array}{l} 0.525 \quad n=22 \\ 0.494 \quad n=23 \end{array} \right\}$$

Ex 14 collisions in hashing



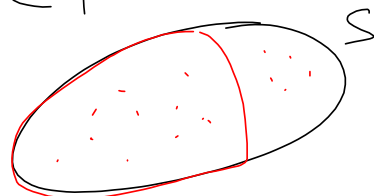
Q: how large does n have to be before $p(\text{collision}) \geq \frac{1}{2}$?

We calculate $q = \text{prob. of no collision}$

$$q = \frac{m}{m} \cdot \frac{m-1}{m} \cdot \dots \cdot \frac{m-n+1}{m} < \frac{1}{2} \text{ when } n \geq 1.77\sqrt{m}$$

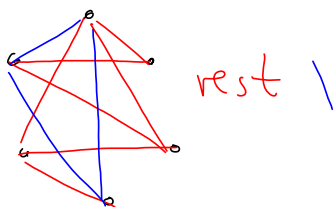
Probabilistic method

basic idea if $P(E) < 1$
then event $\bar{E} \neq \emptyset$



Back to Ramsey theory

recall that \forall 2-colouring of the edges of K_6 there is a monochromatic K_3



define $r(k, k)$ is minimum n s.t.
every 2-col of K_n has a monochromatic K_k

We already know that $r(2, 2) = 2 = 2^{\frac{2}{2}}$ $k=2$
 $r(3, 3) = 6 > 2^{\frac{3}{2}}$

Theorem (Erdős)

$$\forall k \in \mathbb{Z}_+ \quad r(k, k) \geq 2^{\frac{k}{2}}$$

P: \rightarrow assume $n \geq 2^{\frac{k}{2}}$, and $k \geq 4$
consider a random 2-colouring of the edges of K_n
(for each independently: colour it red with $p = \frac{1}{2}$)

We prove that prob that a randomly coloured K_n
has a monochr. K_k is < 1 .

First enumerate all the distinct k -sets of $\{1, 2, \dots, n\}$

$S_1, S_2, \dots, S_{\binom{n}{k}}$ ← each correspond to a unique k_k of K_n

E_i : all edges of S_i have the same colour

$$P(E_i) = \left(\frac{1}{2}\right)^{\binom{k}{2}} \cdot 2$$

$$\bigcup E_i = E_1 \cup E_2 \cup \dots \cup E_{\binom{n}{k}}$$

is the event that at least one k_k is monochr.

$$P(\bigcup E_i) \leq \sum P(E_i)$$

(Union bound)



$$P(\bigcup E_i) \leq \sum_{i=1}^{\binom{n}{k}} P(E_i)$$

$$= 2 \cdot \left(\frac{1}{2}\right)^{\binom{k}{2}} \cdot \binom{n}{k}$$

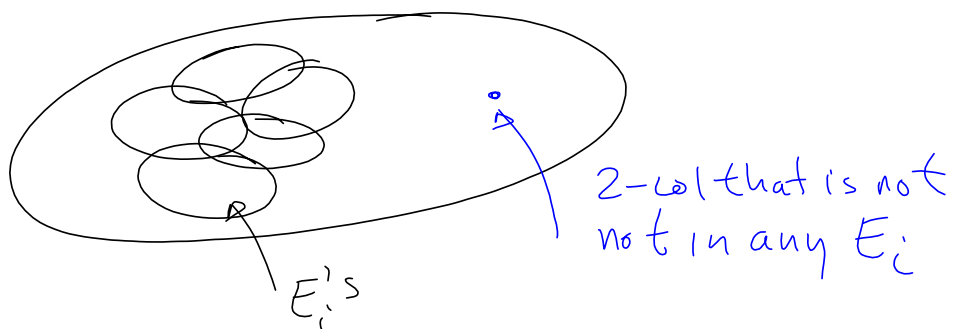
$$\binom{n}{k} < \frac{n^k}{2^{k-1}}$$

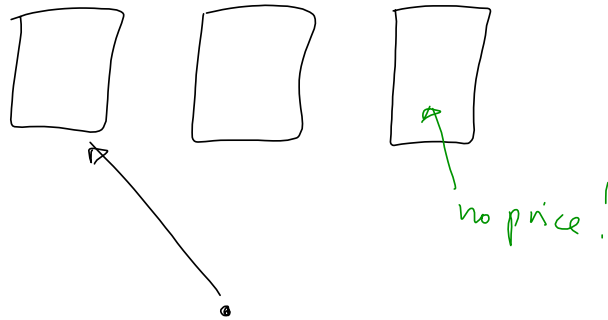
$$n < 2^{k/2}$$

$$< \frac{n^k}{2^{k-1}} \cdot 2 \cdot 2^{-\frac{k}{2}(k-1)}$$

$$< \frac{2^{k^2/2}}{2^{k-1}} \cdot 2 \cdot 2^{-\frac{k}{2}(k-1)} = 2^{2-\frac{k}{2}} \leq 1 \text{ as } k \geq 4$$

Look at the set of all possible
2-colorings of K_n $2^{\binom{n}{2}}$ such





Bayes's Theorem. If E and F are events with $p(E), p(F) > 0$ then

$$p(F|E) = \frac{p(E|F) \cdot p(F)}{p(E|F) \cdot p(F) + p(E|\bar{F}) \cdot p(\bar{F})}$$

P:

By definition $p(F|E) \stackrel{(\Delta)}{=} \frac{p(F \cap E)}{p(E)}$

and $p(F \cap E) = p(E \cap F) \stackrel{(\square)}{=} p(E|F) \cdot p(F)$

we also have

$$\begin{aligned} p(E) &= p(E \cap F) + p(E \cap \bar{F}) \quad \text{as } F \cup \bar{F} = S \\ &\stackrel{(*)}{=} p(E|F) \cdot p(F) + p(E|\bar{F}) \cdot p(\bar{F}) \quad \text{and } F \cap \bar{F} = \emptyset \end{aligned}$$

inserting (\square) and $(*)$ in (Δ) we get

$$\begin{aligned} p(F|E) &= \frac{p(F \cap E)}{p(E)} = \frac{p(E \cap F)}{p(E)} \\ &= \frac{p(E|F) \cdot p(F)}{p(E|F) \cdot p(F) + p(E|\bar{F}) \cdot p(\bar{F})} \end{aligned}$$

