

General Bayes thm.

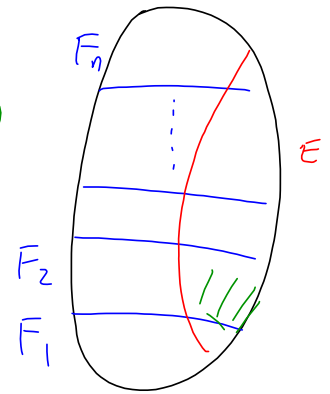
$$\begin{aligned} & \text{by def} \\ & P(A|B) \\ & = \frac{P(A \cap B)}{P(B)} \end{aligned}$$

$$P(F_2|E) = \frac{P(F_2 \cap E)}{P(E)}$$

$$= \frac{P(E|F_2)P(F_2)}{P(E)}$$

$$P(E) = \sum_{i=1}^n P(E|F_i)P(F_i)$$

$$\text{So } P(F_2|E) = \frac{P(E|F_2)P(F_2)}{\sum_{i=1}^n P(E|F_i)P(F_i)}$$



Ex 2 sect 7.3 look at a rare disease

F : person has disease $P(F) = 10^{-5}$

E : positive test

suppose we know the following about the test:

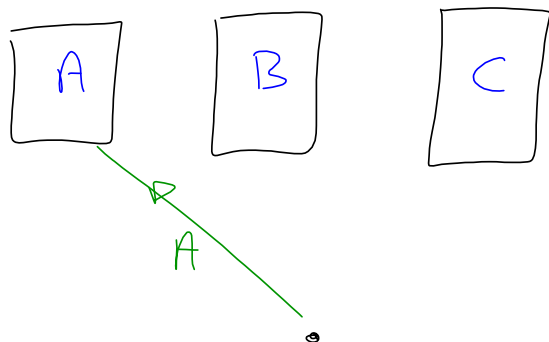
$$P(E|F) = \frac{99}{100}$$

$$P(\bar{E}|\bar{F}) = \frac{995}{1000}$$

We want to know $P(F|E)$ that is, the prob of having the disease if the test is positive.

$$\begin{aligned}
 \left(\begin{array}{c} \bar{F} \\ F \end{array} \right) \in P(F|E) &= \frac{P(E|F)P(F)}{P(E|F)P(F) + P(E|\bar{F})P(\bar{F})} \\
 &= \frac{\frac{99}{100} \cdot 10^{-5}}{\frac{99}{100} \cdot 10^{-5} + \frac{5}{1000} (1 - 10^{-5})} \\
 &\sim 0,002.
 \end{aligned}$$

Monty Hall puzzle



1.

2. host opens door B : no price

3. ? change?

Intuition : we should change to C

Initially we have $p(E_A) = p(E_B) = p(E_C) = \frac{1}{3}$
 where E_A : price behind door A

$$\text{so } p(E_B \cup E_C) = \frac{2}{3}$$

We learned that price is not behind door B
 since host opened door B

$$\text{Thus } p(E_C) = \frac{2}{3} \quad (\text{as now } p(E_B) = 0)$$

analysis using Bayes' formula

$B_{\text{opened}} (B_0)$: the host opens door B

E_A : price behind door A

E_B, E_C similar

we seek $P(E_A | B_0)$ (and $P(E_C | B_0)$)

By Bayes's formula :

$$P(E_A | B_0) = \frac{P(B_0 | E_A)P(E_A)}{P(B_0 | E_A)P(E_A) + P(B_0 | E_B)P(E_B) + P(B_0 | E_C)P(E_C)}$$

$$P(B_0 | E_A) = \frac{1}{2}$$

$$P(B_0 | E_B) = 0 \quad P(B_0 | E_C) = 1$$

$$P(E_A) = P(E_B) = P(E_C) = \frac{1}{3}$$

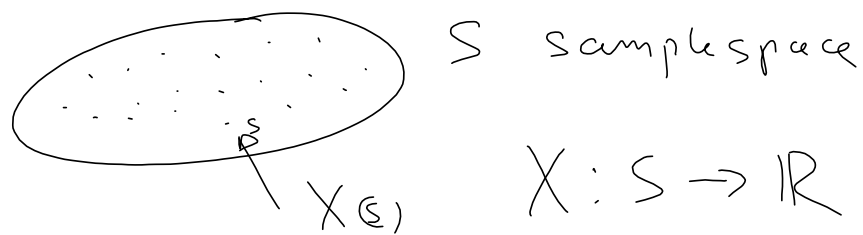
$$\text{so } P(B_0) = \frac{1}{2} \cdot \frac{1}{3} + 0 \cdot \frac{1}{3} + 1 \cdot \frac{1}{3} = \frac{1}{2}$$

$$\text{Now } P(E_A | B_0) = \frac{P(B_0 | E_A)P(E_A)}{P(B_0)}$$

$$= \frac{\frac{1}{2} \cdot \frac{1}{3}}{\frac{1}{2}} = \frac{1}{3}$$

$$P(E_C | B_0) = \frac{1 \cdot \frac{1}{3}}{\frac{1}{2}} = \frac{2}{3} \quad !!$$

Back random variables



ex 2 dice $(i, j) \in \{(1,1) \dots (6,6)\}$
 $X(i, j) = i + j$

Expected value of random var X

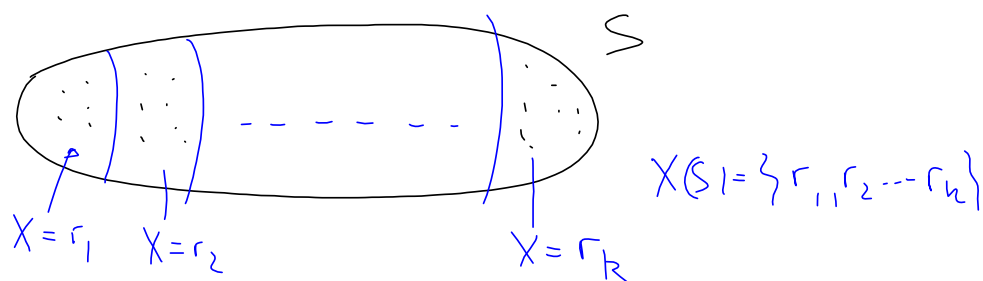
$$E(X) \stackrel{\text{def}}{=} \sum_{s \in S} X(s) \cdot p(s)$$

ex one die
 $X = \text{value}$ $E(X) = \frac{1}{6} (1+2+3+4+5+6)$
 $= \frac{7}{2}$

$$E(X) = \sum_{s \in S} X(s) \cdot p(s)$$

$$= \sum_{r \in X(S)} r \cdot p(X=r)$$

where $X(S)$ is the set of values X can take on elements from S



Linearity of expectation

X_1, X_2, \dots, X_n random variables on S

set $X = X_1 + X_2 + \dots + X_n$

that is $X(\omega) = \sum_{i=1}^n X_i(\omega)$

Theorem $E(X) = E(X_1) + \dots + E(X_n)$

P:

$$E(X) = \sum_{\omega \in S} X(\omega) \cdot P(\omega)$$

$$= \sum_{\omega \in S} (X_1(\omega) + \dots + X_n(\omega)) P(\omega)$$

$$= \sum_{\omega \in S} X_1(\omega) \cdot P(\omega) + \sum_{\omega \in S} X_2(\omega) \cdot P(\omega) + \dots + \sum_{\omega \in S} X_n(\omega) \cdot P(\omega)$$

$$= E(X_1) + \dots + E(X_n)$$

$E(X)$: ^{roll} 2 dice $X_i(\omega) =$ value of dice i $i=1,2$

$X = X_1 + X_2$ (sum of values)

Find $E(X)$

method 1:

$$\begin{aligned} E(X) &= \sum_{r \in X(\Omega)} r \cdot p(X=r) \\ &= 2 \cdot p(X=2) + 3 \cdot p(X=3) + \dots \\ &\quad + 12 \cdot p(X=12) \end{aligned}$$

method 2 (use linearity of expectation)

$$\begin{aligned} E(X) &= E(X_1) + E(X_2) \\ &= \frac{7}{2} + \frac{7}{2} = 7 \end{aligned}$$

(since $E(X_i) = \frac{1}{6}(1+2+3+4+5+6) = \frac{7}{2}$)

what if we roll 100 dice or 1000

still easy with method 2 (contrary to method 1)

because if $n=100$:

$$E(X) = 100 \cdot \frac{7}{2} = 350$$

Note that we also have

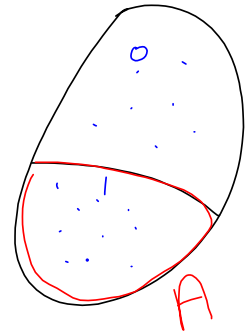
if Y is a random var on S
and a, b are constants, then

$$E(aY+b) = aE(Y)+b$$

Indicator random variables

Let A be an event in sample space S
and define the random var X_A as

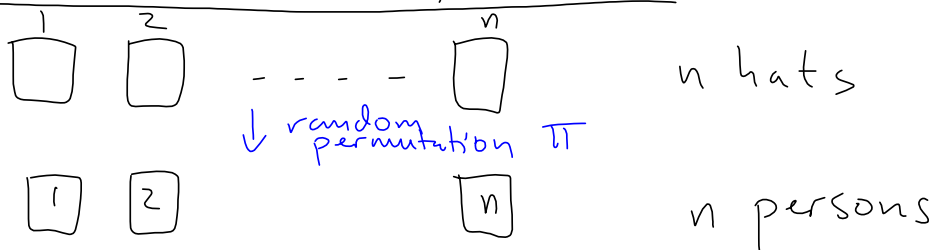
$$X_A(s) = \begin{cases} 1 & \text{if } s \in A \\ 0 & \text{if } s \notin A \end{cases}$$



$$\text{so } P(X_A=1) = P(A) = p$$

$$\text{and (!)} E(X_A) = 1 \cdot p + 0 \cdot (1-p) = p$$

Ex 6 Hatcheck problem



E_i : person i gets his hat ($P(E_i) = \frac{1}{n}$)

$$X_i = X_{E_i} = \begin{cases} 1 & \text{if } i \text{ gets his hat } p = \frac{1}{n} \\ 0 & \text{otherwise} \end{cases}$$

Define X as

$$X = X_1 + X_2 + \dots + X_n$$

then X is precisely # of people who get their own hat

$$\text{so } E(X) = \sum_{i=1}^n E(X_i) = \sum_{i=1}^n \frac{1}{n} = n \frac{1}{n} = 1$$

Thm 2 Let $X = \#$ successes in n independent Bernoulli expts with success prob p .

$$\text{Then } E(X) = np$$

recall p (exactly k success in n expts) is given by $b(k, n, p) = \binom{n}{k} \cdot p^k (1-p)^{n-k}$

$$\text{so } E(X) = \sum_{k=0}^n k \cdot b(k, n, p) = np \quad \uparrow \text{ we claim}$$

Define $X_i = \begin{cases} 1 & \text{if exp } i \text{ is success } P(X_i=1)=p \\ 0 & \text{otherwise} \end{cases}$

Then $X = \sum_{i=1}^n X_i$ is # of successes

$$\text{and } E(X) = E\left(\sum_{i=1}^n X_i\right) = \sum_{i=1}^n E(X_i) = \sum_{i=1}^n p = np$$

Geometric distribution

X has a geometric distribution with prob p if $P(X=k) = (1-p)^{k-1} \cdot p$

$$E(X) = \sum_{j=1}^{\infty} j \cdot P(X=j)$$

$$= \sum_{j=1}^{\infty} j \cdot (1-p)^{j-1} \cdot p$$

$$= p \sum_{j=1}^{\infty} j \cdot (1-p)^{j-1} = p \cdot \frac{1}{p^2} = \frac{1}{p}$$

note $\sum_{j=1}^{\infty} (1-p)^{j-1} p = p \sum_{j=1}^{\infty} (1-p)^{j-1} = p \frac{1}{(1-(1-p))} = p \cdot \frac{1}{p} = 1$

